Prevalence of Congenital Anomalies in the Japan Environment and Children’s Study

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ABSTRACT

Background: The aims of the present report were to estimate the prevalence of congenital anomalies (CAs) among infants in Japan using data from the Japan Environment and Children’s Study (JECS) and to evaluate the validity of CA classification within JECS.

Methods: Data on CAs were collected at delivery and at age 1 month from the medical records of 101,825 infants at 15 regional centers. The analyses focused on 61 CAs, selected on the basis of reported associations with environmental exposure. Prevalence per 10,000 pregnancies (including miscarriages, stillbirths, and live births) was stratified according to four reporting patterns (at delivery, at age 1 month, at either, and at both). To evaluate the accuracy of observed CA prevalence, the medical records of 179 cases from a single JECS regional center underwent independent, retrospective re-evaluation.

Results: The prevalence of major CAs in four reporting patterns (at delivery, at age 1 month, at either, and at both) was 2.4, 2.6, 3.5, and 1.4 for myelomeningocele/spina bifida; 4.3, 4.2, 5.3, and 3.2 for cleft palate; 18.1, 17.4, 19.5, and 15.1 for cleft lip with or without cleft palate; 73.4, 100.3, 120.8, and 52.8 for congenital heart disease; and 10.5, 14.1, 15.0, and 9.6 for Down’s syndrome, respectively. In the subsample re-evaluation, CA diagnoses were confirmed for 92.7%, 93.3%, 90.5%, and 97.8% of cases in the four reporting patterns (at delivery, at age 1 month, at either, and at both), respectively.

Conclusions: The present report generated reliable data concerning the prevalence of major CAs in JECS.

Key words: congenital anomalies; Japan Environment and Children’s Study; JECS; prevalence; birth cohort

INTRODUCTION

Congenital anomalies (CAs) are an important cause of not only infant and childhood deaths but also chronic illness and lifelong disability. To date, approximately 50% of CAs have been linked to a specific cause, genetic, socioeconomic and demographic, environmental, infectious, and maternal nutritional risk factors are known at present. Environmental toxins that are known to cause several CAs are one of the risk factors that can be minimized by appropriate standards and regulations. The reported risk factors for congenital heart disease include prenatal exposure to pesticides, organic solvents, and air pollution. Exposure to some endocrine disrupters has been linked to cryptorchidism and hypospadias. Pesticide exposure is associated with urinary malformedness. However, evidence for many of these relationships is still limited.

The Japan Environment and Children’s Study (JECS) is a nationwide prospective birth cohort study conducted by the Ministry of the Environment of Japan. The aims of JECS are to identify environmental factors, in particular chemical substances, with a detrimental impact on child health and development, and to facilitate improved environmental risk management system. CAs are one of the priority health outcomes of JECS. However, the collection of CA data in JECS is hampered by the fact that, to date, no population-based CA registration or surveillance system has been established in Japan. At its 15 regional centers, JECS investigators therefore collect data on CAs from the medical records of infants at delivery and at 1 month of age.

The aims of the present report were to estimate the prevalence of CAs among infants from Japan using data from the JECS cohort and to evaluate the validity of CA classification within JECS.
METHODS

Study design
Recruitment of women in the first trimester of pregnancy took place between January 2011 and March 2014 at the 15 JECS regional centers. The respective offspring will undergo follow-up until at the age of 13 years. JECS protocols are described in detail elsewhere.\(^7\)\(^8\) During the JECS recruitment period, a total of 104,102 infants were enrolled through hospitals or local government offices. In the event of miscarriage, medical termination of pregnancy, stillbirth, or neonatal mortality, the respective regional center collected and recorded the necessary data prior to termination of JECS participation. All types of obstetric facilities (perinatal medical centers, maternity hospitals, and maternity homes) cooperated with JECS to increase the participation rate. No data concerning the participation rate across the entire JECS recruitment period were collected, since the recruitment area was expanded in 2012 and recruitment during the entire JECS recruitment period were collected, since the participation rate in 2011 and 2014 was restricted to a period of 3 months. However, in 2013, when recruitment had largely stabilized, JECS collected data on 45% of all total live births across its 15 geographical recruitment regions.\(^7\)

All JECS procedures adhere to the Ethical Guidelines for Epidemiological Research of the Japan Ministry of Health, Labour, and Welfare. JECS protocols were reviewed and approved by the Institutional Review Board of the Ministry of the Environment, and by the ethics committees of all participating institutions. Written informed consent was obtained from all pregnant women prior to inclusion.

Cohort used for the investigation of CA prevalence
For the purpose of the present report, the following JECS cases were excluded: (i) cases in which no data at the CA reporting time-points were available (\(n = 2,277\)), and ii) cases in which pregnancy outcome was unknown (\(n = 38\)).

Some participants were withdrawn from the study after the completion of data collection at delivery due to miscarriage (\(n = 1,340\)) and stillbirth (\(n = 291\)). Follow-up of these participants was terminated following the collection of data at delivery. In 705 participants, data were available at the delivery time-point but not at 1 month. In the case of 44 participants, data were available at the 1 month time-point but not at delivery. In each case, the reason for this was undocumented.

Collection of congenital anomaly data
The present analyses focused on a list of 61 CAs (Table 1). This represents a modified version of the list used in the Hokkaido Study on Environment and Children’s Health. The latter comprised 55 CAs, which were selected as possible markers of environmental exposure.\(^9\)

CA data were collected by transcribing information entered in the infants’ medical records at delivery and at 1 month of age onto JECS transcription forms. This information included the respective 10\(^{th}\) revision of the International Statistical Classification of Diseases and Related Health Problems (ICD) code (Table 1), and the terminology used to denote the CA in the medical records. In each case, this process was completed by a physician, a midwife, a nurse, or a trained research coordinator. Each CA was listed in the transcription form in a separate box.

Categorization of anomalies described in the comment section
When CAs not listed in Table 1 were found or suspected in medical records, the information was transcribed into a comment section (ie, uncategorized CAs), to reduce the number of cases overlooked by the person making the transcription.

In total, 13,144 uncategorized CAs were recorded in the comment sections. Two pediatricians (H.M. and Y.Y.) independently evaluated the respective comments and categorized the described anomalies as specific CAs. If uncategorized CA information containing misspellings, unexplained abbreviations, or any ambiguity were detected, these were corrected by the pediatricians prior to categorization. For the purposes of the present analyses, suspected CAs were assessed using the comment sections of the transcription forms only. The respective medical records were not consulted. Suspected CAs still awaiting formal diagnosis were not categorized as CAs for the purposes of the present analyses. In cases of uncertainty, the two pediatricians assigned a consensus diagnosis.

For the valuation of entries in the comments section, 31 CAs that were easily detected at delivery and that generally required prompt medical attention after delivery were selected from the list of 61 CAs of interest. The 31 selected CAs comprised anencephaly, encephalocele, hydrocephaly, holoprosencephaly, ablepharon, microphthalmia/anophthalmia, congenital cataract, cleft lip (CL), cleft palate (CP), cleft lip and palate (CLP), facial cleft, diaphragmatic hernia, congenital heart disease, omphalocele, gastrochisis, esophageal atresia with or without fistula, duodenal atresia/stenosis, intestinal atresia/stenosis, anorectal atresia/stenosis, hypospadias, undescended testis/cryptorchidism, polydactyly of the fingers, syndactyly of the fingers, cleft hand, polydactyly of the toes, syndactyly of the toes, cleft foot, myelomeningocele/spina bifida, Down syndrome, Trisomy 18, and Trisomy 13. Through this process, a total of 755 uncategorized CAs (5.74%) were categorized as CAs.

Covariates
Data were also collected on mother’s age at delivery, regional center, singleton/multiple pregnancy, and pregnancy outcomes (live birth, spontaneous abortion, medical termination of pregnancy, and stillbirth). Stillbirth was defined as the occurrence of fetal death at >21 weeks gestation.

Analysis of CA prevalence
Data collected at delivery only were stratified according to four pregnancy outcomes: 1) spontaneous abortion, 2) medical termination of pregnancy, 3) stillbirth, and 4) live birth. All cases in which CAs had been recorded at the 1 month time-point only were classified as live births.

In a first step, data collected both at delivery and at 1 month were stratified according to the four possible CA reporting patterns: 1) CA recorded at delivery time-point only, 2) CA recorded at 1 month time-point only, 3) CA recorded either at delivery or at 1 month, and 4) CA recorded at both delivery and the 1 month time-point. For each of these four categories, the type, number, and percentage of CAs are shown in Table 1. A high CA prevalence was observed in the following two data collection patterns: 1) data collected at delivery only, and 2) data collected at 1 month of age only. The number of CAs recorded in these two data collection patterns are shown in Table 2.
Table 1. Number and percentage of the 61 congenital anomalies of interest according to reporting and data collection patterns

<table>
<thead>
<tr>
<th>ICD-10 codec</th>
<th>Data collected at delivery only</th>
<th>Data collected at one month of age only</th>
<th>Data collected both at delivery and at one month of age</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>N</td>
<td>1,010</td>
<td>330</td>
</tr>
<tr>
<td>Anencephaly</td>
<td>Q00</td>
<td>3</td>
<td>0.297</td>
</tr>
<tr>
<td>Encephalocele</td>
<td>Q01</td>
<td>1</td>
<td>0.099</td>
</tr>
<tr>
<td>Microcephaly</td>
<td>Q02</td>
<td>0</td>
<td>0.000</td>
</tr>
<tr>
<td>Hydrocephaly</td>
<td>Q03</td>
<td>1</td>
<td>0.099</td>
</tr>
<tr>
<td>Craniosynostosis</td>
<td>P96.3</td>
<td>0</td>
<td>0.000</td>
</tr>
<tr>
<td>Holoprosencephaly</td>
<td>Q04.2</td>
<td>2</td>
<td>0.198</td>
</tr>
<tr>
<td>Agenesis of corpus callosum</td>
<td>Q04.0</td>
<td>0</td>
<td>0.000</td>
</tr>
<tr>
<td>Anophthalmos/microphthalmos</td>
<td>Q11.0–Q11.2</td>
<td>0</td>
<td>0.000</td>
</tr>
</tbody>
</table>

**Notes:**
- a Each congenital anomaly is listed with its number of cases and percentage per data collection patterns.
- b Developmental dysplasia of the hip was included in the transcription form at 1 month of age only.
- c Developmental dysplasia of the hip was included in the transcription form at 1 month of age only.
- dTranscripts of both the ICD-10 code and congenital anomaly name in medical records were collected at the data collection phase.
In a second step, data from the two data collection patterns showing a high prevalence of CAs (delivery only and 1 month of age only) were stratified for the four CA reporting patterns: 1) CA reporting at delivery data collection only (D-DELIVERY), 2) CA reporting at 1 month data collection only (D-1M), 3) CA reporting either at delivery or at 1 month data collection (D-OR), and 4) CA reporting both at delivery and at 1 month data collection (D-AND). For each of the four reporting patterns, the prevalence of the CA per 10,000 pregnancies was then calculated.

Categorization of CP, CL, and CLP

In accordance with other epidemiological studies, CL and CLP were re-categorized as ‘cleft lip with or without cleft palate’ (CL/P). In eight cases, CLP was reported at one time and CL at another time. In four cases, CL and CLP were reported at the same time. In one case, CP and CL were reported at the same time. Therefore, 13 cases were recorded as “not classified as CL, CP, or CLP” in Table 1. Of these, eight cases were categorized as CL/P for the purposes of the present analyses.

### Table 2. Number and prevalence per 10,000 pregnancies of the 61 congenital anomalies of interest according to reporting pattern

<table>
<thead>
<tr>
<th>Congenital anomaly</th>
<th>Total number</th>
<th>Prevalence per 10,000 pregnancies</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>D-DELIVERYa</td>
<td>D-1Mb</td>
</tr>
<tr>
<td>Total number</td>
<td>101,825</td>
<td>101,825</td>
</tr>
<tr>
<td>Major anomalies</td>
<td>1,959</td>
<td>2,437</td>
</tr>
<tr>
<td>Central nervous system</td>
<td>174</td>
<td>228</td>
</tr>
<tr>
<td>Neural tube defect</td>
<td>65</td>
<td>61</td>
</tr>
<tr>
<td>Anencephaly</td>
<td>28</td>
<td>21</td>
</tr>
<tr>
<td>Encephalocele</td>
<td>13</td>
<td>15</td>
</tr>
<tr>
<td>Myelomeningocele/Spina bifida</td>
<td>24</td>
<td>26</td>
</tr>
<tr>
<td>Hydrocephaly</td>
<td>43</td>
<td>72</td>
</tr>
<tr>
<td>Microcephaly</td>
<td>11</td>
<td>29</td>
</tr>
<tr>
<td>Holoprosencephaly</td>
<td>23</td>
<td>38</td>
</tr>
<tr>
<td>Cranioptases</td>
<td>35</td>
<td>35</td>
</tr>
<tr>
<td>Age of corpus callosum</td>
<td>17</td>
<td>19</td>
</tr>
<tr>
<td>Eye</td>
<td>33</td>
<td>24</td>
</tr>
<tr>
<td>Anophthalmos/microphthalmos</td>
<td>16</td>
<td>12</td>
</tr>
<tr>
<td>Congenital cataract</td>
<td>17</td>
<td>13</td>
</tr>
<tr>
<td>Oro-facial cleft</td>
<td>232</td>
<td>223</td>
</tr>
<tr>
<td>Cleft palate</td>
<td>44</td>
<td>43</td>
</tr>
<tr>
<td>Cleft lip with or without palate</td>
<td>184</td>
<td>177</td>
</tr>
<tr>
<td>Facial cleft</td>
<td>7</td>
<td>4</td>
</tr>
<tr>
<td>Respiratory system</td>
<td>35</td>
<td>40</td>
</tr>
<tr>
<td>Intralobar sequestration</td>
<td>1</td>
<td>4</td>
</tr>
<tr>
<td>Congenital cystic adenomatoid malformation</td>
<td>10</td>
<td>12</td>
</tr>
<tr>
<td>Pulmonary hypoplasia</td>
<td>24</td>
<td>26</td>
</tr>
<tr>
<td>Congenital heart disease</td>
<td>747</td>
<td>1,021</td>
</tr>
<tr>
<td>Cardiac arrhythmia</td>
<td>82</td>
<td>50</td>
</tr>
<tr>
<td>Abdominal wall defects</td>
<td>61</td>
<td>286</td>
</tr>
<tr>
<td>Omphalocele</td>
<td>43</td>
<td>272</td>
</tr>
<tr>
<td>Gastrochisis</td>
<td>20</td>
<td>16</td>
</tr>
<tr>
<td>Digestive system</td>
<td>121</td>
<td>100</td>
</tr>
<tr>
<td>Esophageal atresia with or without fistula</td>
<td>19</td>
<td>19</td>
</tr>
<tr>
<td>Duodenal atresia/stenosis</td>
<td>17</td>
<td>13</td>
</tr>
<tr>
<td>Intestinal atresia/stenosis</td>
<td>16</td>
<td>10</td>
</tr>
<tr>
<td>Anorectal atresia/stenosis</td>
<td>39</td>
<td>37</td>
</tr>
<tr>
<td>Diaphragmatic hernia</td>
<td>34</td>
<td>28</td>
</tr>
</tbody>
</table>

In cases in which information on the CA could not be located in the medical or insurance records, the physicians categorized the anomaly as ‘not classified as CL, CP, or CLP’.
the transcription form as erroneous and added a comment to the data. The re-evaluation results were categorized as: 1) confirmed or 2) not confirmed. The number and percentage of CAs in each category were then calculated according to the data reporting pattern.

### Statistical analysis

All analyses were performed using SAS version 9.4 (SAS Institute, Cary, NC, USA). Binary variables are shown as the number of persons and percentages (%) by category. Incidence is the total number of each CA case in the data reporting pattern. CA cases categorized according to organ were counted if any subordinate CAs were reported. In this previous report, prevalence estimated per 10,000 pregnancies (including spontaneous abortion, medical termination of pregnancy, stillbirth, and live birth) was stratified according to the reporting pattern. All prevalence included any type of data collection.

### RESULTS

The final analyses included data from a total of 101,825 infants. Figure 1 shows the Venn diagram of the three data collection patterns: (i) both at delivery and at 1 month of age (n = 99,445 infants), (ii) at delivery only (n = 2,336 infants), and (iii) at 1
month of age only (n = 44 infants). Pregnancy outcomes in the “at delivery only group” comprised 291 stillbirths (0.29%), 1,010 spontaneous abortions (0.99%), 330 medical terminations of pregnancies (0.32%), and 705 livebirths (0.69%). The numbers of infants at each regional center, stratified according to maternal age at delivery, are shown in Table 1.

Table 1 shows the number and percentages of 61 CAs of interest, stratified according to reporting pattern. In the spontaneous abortion, medical termination of pregnancy, and stillbirth groups, several major CAs were reported. In particular, nine Down syndrome cases (2.73%) were found in the medical termination of pregnancy group, while six Down syndrome cases (0.85%) were reported in the live birth group.

Among live births in the group with at delivery data only, major CAs were reported in 705 infants. Two major CAs were also reported in the group with 1 month data only. Among data collected both at delivery and at 1 month, the total number of each CA varied according to data reporting pattern, with the exception of amniotic band constriction. The number of omphalocele cases was substantially higher at 1 month than at delivery.

Prevalence of each congenital anomaly
Table 2 summarizes the number and prevalence per 10,000 pregnancies of the 61 CAs of interest, stratified according to reporting pattern. The prevalence of major anomalies in each reporting pattern ranged from 126.7 to 298.6 per 10,000 pregnancies. D-OR had the highest and D-AND the lowest prevalence. Compared with D-DELIVERY and D-1M, the prevalence of almost all major CAs was similar or higher in D-1M. Concerning prevalence according to organ, congenital heart disease showed the highest prevalence, followed by ~20 cases per 10,000 pregnancies for limb, urinary system, oro-facial cleft, and central nervous system. By contrast, the prevalence of eye, ear, respiratory system, genital system, and skeletal dysplasia CAs was low, at less than 10 per 10,000 pregnancies. The prevalence of digestive system and chromosomal CAs was 10 per 10,000 pregnancies. There was a difference in the number of abdominal wall defects between D-DELIVERY (6.0) and D-1M (28.1).

Independent evaluation of congenital anomaly data accuracy
Of the 179 CA cases that underwent re-evaluation, one was found to have been miscoded in the transcription form, eight were withdrawn from the study after delivery, and two were not recorded due to the hospital’s refusal to disclose the information. Details of the remaining 168 cases are shown in Table 3. In the D-DELIVERY group, 115 CAs (92.7%) were confirmed and nine (7.3%) were not confirmed. In the D-1M group, 126 CAs (93.3%) were confirmed and nine (6.7%) were not confirmed. In the D-OR group, 152 CAs (90.5%) were confirmed and 16 (9.5%) were not confirmed. In the D-AND group, 89 CAs (97.8%) were confirmed and two (2.2%) were not confirmed. For 16 cases, no description

Prevalence of Congenital Anomalies in JECS
of the CA was found in the respective medical or and health insurance records. The 16 suspected CAs comprised diaphragmatic hernia (n = 5), congenital heart disease (n = 3), congenital hydronephrosis (n = 2), duodenal atresia/stenosis (n = 1), intestinal atresia/stenosis (n = 1), myelomeningocele/spina bifida (n = 1), omphalocele (n = 1), polydactyly of the fingers (n = 1), and undescended testis/cryptorchidism (n = 1). Of the five cases of unreported diaphragmatic hernia, four were in the D-DELIVERY and one was in the D-1M. There were no other CAs without descriptions in the medical and insurance records in the non-conformed 16 cases. eTable 2 shows the number and proportion of CAs for which confirmation in accordance with the ICD-10 code entered in the respective medical records was achieved.

**DISCUSSION**

Using data from the JECS nationwide prospective birth cohort study, the present report analyzed: (i) data collection patterns, and (ii) the number and prevalence of 61 CAs of interest, which were stratified according to the pattern of data reporting. A high CA prevalence was observed in data collected at a single time-point (at delivery only or at 1 month of age only). In addition, differing numbers of CAs were observed between the different data reporting patterns. To investigate this, an independent evaluation was performed of the accuracy of the transcribing of data from medical records to the study transcription forms across each of the data reporting patterns. This was performed using a subsample from the Kochi regional JECS center. In this evaluation, confirmation in the medical/insurance records was achieved for 90.5% of CAs in the D-OR and 97.8% of CAs in the D-AND. We also reported differences in prevalence according to organ.

In the present report, at least 90.5% of the CAs reported in the transcription forms were confirmed after re-evaluation of the corresponding medical and insurance records. Transcription errors were mainly due to misclassifications that caused a bias toward the null value when the misclassification rate was equal between the two compared groups.9 In the JECS protocol, the misclassification was equal for exposures to environmental toxicants because the assessors were blinded to the exposure measurements. Rui et al investigated the impact of misclassification and measurement errors using a simulation study approach. In their study, when the prevalence of one disease was low (3.2%) and the effect size was small (0.3), a power of 15.1% was observed between no misclassification and high sensitivity and specificity scenario (sensitivity 0.84 and specificity 0.96) with a small amount of covariate measurement error.10 By contrast, the JECS sample size was calculated with sufficient power. For instance, to test a hypothesis concerning a disorder with a prevalence of 0.1%, such as Down syndrome, with a relative risk of 2.0 and an alpha error of 0.05, and using a cohort in which the proportion of individuals with a high level of exposure to the chemical substance of interest is 25%, a sample size of 64,536 participants is required to provide a statistical power of 80%.11 The JECS cohort exceeds this sample size by 101,825. Thus, even if around 15% of the statistical power was lost, the JECS CA data still retained sufficient power to address the study hypotheses. Thus, since data accuracy in the Kochi report had a sensitivity of >0.84 and a specificity of >0.96 during the transcription procedure, the JECS had sufficient accuracy in reporting CAs.

In the data accuracy report of the D-OR group, the Kochi regional center enrolled a total of 7,140 infants. Of these, 168 (2.4%) reported CAs and 6,972 (97.6%) did not. In the reported CA group, 152 (2.1%) were confirmed and 16 (0.2%) were not. Because of the high number of non-reported CA infants, our report could not provide an accuracy rate for non-reported infants. However, other mistakable CAs diagnosis in respective medical and insurance record were not found among non-conformed cases. This might suggest a risk of mistakes in using the transcription forms. The number of transcription errors for non-reported CAs infants was expected to be at least similar to the number of non-conformed CAs (16 cases). Using a sensitivity of >0.84 and a specificity of >0.96, the number of transcription errors in cases in the non-reported group was calculated to be 28 (sensitivity 0.84 and specificity 0.99). This number was more than expected. Taking this into consideration, the JECS CA data can be considered to have provided acceptable power for detecting the impact of environmental factors on the CA occurrence. By contrast, while D-OR had the highest number of CAs, it also had the highest rate of misclassification. This misclassification would be expected to cause an underestimation or an overestimation of CA prevalence. eTable 2 provides evidence of overestimation from the ICD-10 code. However, infants born in maternal homes were not diagnosed with any CAs using the ICD-10 code because midwives are not qualified to use the ICD-10 code in Japan. For these reasons, the present authors recommend the use of two different CA reporting groups in future analyses of CA prevalence, for example, D-1M and D-OR.

In the medical termination of pregnancy and stillbirth groups, a high proportion of life-threatening CAs and chromosomal diseases were found. Schechtman et al reported a direct correlation between CA severity and the rate of medical termination of pregnancy rates.12 Furthermore, a study in Japan showed that medical termination of pregnancy was performed in 93.3% of cases in which a chromosomal anomaly was confirmed via prenatal conventional cytogenetic analysis at 15–18 weeks of pregnancy.13 However, in the present report, life-threatening CAs were also recorded in the spontaneous abortion group. These comprised three cases of anencephaly, three cases of gastroschisis, and one case of an omphalocele. Severe CAs were also recorded in live births from the group with delivery data only. These included CLP, hydrocephaly, diaphragmatic hernia, and esophageal atresia with or without fistula. Among the 44 cases for whom data were collected at 1 month of age only, two CAs were recorded. These results suggest that JECS participants with only a single data collection time-point (ie, at delivery only or at 1 month only) may have a high CA risk. To account for this, CA prevalence was determined using data collected both at delivery and at 1 month.

To describe the CA data characteristics, we compared the CAs in our previous study with those in other registrations. Previous estimates of CP and CL/P prevalence per 10,000 births were 5.83 and 22.76, respectively, in the 2012 report of the Japanese Association of Obstetricians and Gynecologists (JAOG)14; 5.88 and 8.26, respectively, in the 2011–2015 survey of the European Surveillance of Congenital Anomalies (EUROCATS)15; and 6.35 and 10.63, respectively, in the 2004–2006 survey of the National Birth Defects Prevention Network (NBDPN) in the United States.16 CP prevalence in each of these three reports was similar to, or slightly higher than, that found in the present report. Compared to the present report, CL/P prevalence in EUROCATS
and NBDPN were lower, while CL/P prevalence in JAOG was higher. For spina bifida, the prevalence in the present analyses was lower than the prevalence reported in JAOG (5.18) and EUROCATS (5.11) and similar to or lower than the prevalence reported in NBDPN (3.50).

These cross-study similarities and differences in the prevalence of these major CAs are attributable to three factors. First, the findings are influenced by selection bias. The research facilities of the JAOG are mainly based at hospitals that provide care for women with high-risk pregnancies, whereas the JECS collects data from pregnant women who are managed in general hospitals and maternity homes. Since spina bifida requires surgery immediately after delivery, women with spina bifida pregnancies are usually referred to tertiary centers. By contrast, surgery for congenital heart disease is usually performed twice, at around 12 and 20 weeks gestational age, and once after 36 weeks. In the United Kingdom in 2012. As expected, approximately 70% of CL/P cases were diagnosed before 1 month of age.38 These reports showed an increased rate of prenatal CA diagnosis as a result of prenatal screening. In Japan, the guidance given to pregnant women at regular health checks recommends testing by ultrasound at least four times: twice before 23 weeks gestational age, once between 24 and 35 weeks, and once after 36 weeks. In the United Kingdom and Australia, prenatal ultrasound screening is performed twice, at around 12 and 20 weeks gestational age. The CA detection rate of JECS was at least as high as that of the United Kingdom in 2012. As expected, approximately 70–80% of CAs cases in JECS were diagnosed before the age of 1 month.

The present report has three main strengths. First, JECS CA data are based on reports from hospitals and maternity clinics throughout Japan. Reporting bias was not considered, since data collection was performed by research facility personnel rather than by parents. Second, the rate of loss to follow-up was low (2.2%). Third, CA prevalence was estimated on the basis of all pregnancies in the cohort, including all cases of spontaneous abortion and medical termination of pregnancy. Such data, in particular those concerning spontaneous abortion, are not recorded in other CA registries.15,16,37,38 However, these data are important in terms of the identification of the environmental causes of CAs.

The present report also had four limitations that warrant mention. First, only the 61 CAs of interest were considered. Second, misclassification may have occurred due to the
mistranscription of medical records, which could have caused an overestimation or underestimation of CA prevalence. Third, JECS did not collect data from alternative sources, such as child health systems or departments specializing in the diagnosis of CAs after 1 month of age. This is likely to have led to unidentified CA cases being ignored after 1 month of age. To identify overlooked CA cases, the JECS consortium plan to collect additional CA data on 2-year-old children receiving hospital treatment through the use of parent-report questionnaires. Finally, methodological differences between the JECS and other registries may have resulted in the differing estimates of CA prevalence.30 Once the JECS data become open access, the present authors recommend that two data reporting groups, for example, D-1M and D-OR, are used for the analysis of each CA. The number of omphalocele recorded at the 1 month data collection time-point may have been an overestimate, due to misclassification secondary to the similar Japanese terminology for omphalocele and umbilical hernia. Thus, for omphalocele, D-DELIVERY or D-AND data should be used to determine prevalence. There were some instances of misclassification in the JECS data. Since JECS is not a population-based registration program, the data cannot be used to generate definite CA prevalence rates. However, the demographic characteristics of the JECS pregnant women cohort are almost identical to those reported in a previous national survey.7

Conclusions
The present report generated reliable data concerning the prevalence of major CAs in the JECS cohort.

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Authors’ contributions: Research staff at the 15 regional centers of the JECS collected data. The final version of the dataset (jecs-ag-20160424) was fixed and released by the National Center of the JECS (The National Institute for Environmental Studies). AT conducted statistical analysis of the data set. All authors contributed to analysis of the study results. HM wrote the first draft of the manuscript. All authors approved the final version of the manuscript.

APPENDIX A. SUPPLEMENTARY DATA
Supplementary data related to this article can be found at https://doi.org/10.2188/jcea.JE20180014.

APPENDIX B.
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